15th International Conference Applications of Computer Algebra

Teaching Principal Components Analysis with Minitab

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Introduction

The purpose of this presentation is to introduce the logical and arithmetic operators and simple matrix functions of Minitab[®] –a wellknown software package for teaching statistics- as a computer-aid to teach **Principal Components** Analysis (PCA) to graduate students in the field of Education.



PCA, originally proposed by Pearson (1901) is a mathematical technique –a vector space transform- that has its roots in linear algebra and in statistics.

Its main purpose is to reduce a correlated multidimensional data set to an uncorrelated lower dimensional space with maximum variance.

PCA concepts can be a roadblock for non-mathematical oriented students, since statistical definitions (i.e., variance-covariance, correlation) need to be connected to matrix algebra (eigenvectors of a variance-covariance matrix) and to graphical vector representation (including matrix rotation).





 Given m points in a n dimensional space, for large n, how does one project on to a low dimensional space while preserving broad trends in the data and allowing it to be visualized?

Choose a line that fits the data so the points are spread





A sample of *n* observations in the 2-D space



Goal: to account for the variation in a sample in as few variables as possible, to some accuracy

Formally, minimize sum of squares of distances to the line.



Why sum of squares? Because it allows fast minimization, assuming the line passes through 0

Minimizing sum of squares of distances to the line is the same as maximizing the sum of squares of the projections on that line, thanks to Pythagoras.





Y1 and Y2 are new coordinates.

Y1 represents the direction where the data values have the largest uncertainty.Y2 is perpendicular to Y1.

To find Y1 and Y2, we need to make transformation from X1 and X2. To simplify the discussion, we move the origin to (\bar{x}_1, \bar{x}_2) and redefine the (X1,X2) coordinate as

x1 = X1 - $\overline{x_1}$, x2 = X2 - $\overline{x_2}$, so that the origin is (0,0).

The relationship is illustrated in the following graph. We would like to present the data of a given lab, p = (x1,x2) in terms of p = (y1,y2).



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16	13	6	0.47080	-0.95411	-2.0659	-14.1721	11.9346	-0.46098								
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18	13	15	0.47080	0.88072	-10.6963	-19.0682	19.5927	7.58499								
19	13	17	0.47080	1.28846	-12.6141	-20.1563	21.2945	9.37299								
20	14	7	0.69288	-0.75024	-2.7412	-15.5551	13.3108	-0.01505								
21	15	13	0.91495	0.47298	-8.2111	-19.6583	18.9416	4.90085								
22	17	13	1.35911	0.47298	-7.6438	-21.3365	19.9922	4.00471								
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For any given value of theta, then, it is a simple matter to work out the values of Y1 for each of our twenty observations. When 0 is 5 degrees, for example, the calculations are:

		Variable
Z1	Z2	Y1
1.90	0.47	1.94
0.99	0.85	1.06
1.22	0.09	1.22
0.54	-0.68	0.47
0.31	0.47	0.35
0.08	-1.25	-0.03
-0.15	-0.30	-0.17
-0.60	0.09	-0.59
-1.06	-1.63	-1.20
-1.29	-1.25	-1.39
-1.74	-1.63	-1.88
-1.52	-1.06	-1.60
0.76	0.47	0.80
1.90	2.57	2.12
0.31	0.28	0.33
-0.38	-0.10	-0.38
-0.15	0.66	-0.09
-0.15	0.85	-0.07
-0.38	0.66	-0.32
-0.60	0.47	-0.56

```
Mean (X1) = 7.65; VAR (X1) = 19.23
Mean (Z1) = Mean (Z2) = 0;
Variance (Z1) = Variance (Z2) = 1
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VAR (Y1) = $(\cos\theta) x1 + (\sin\theta) x2 = 1.12$

Note that each of the original variables has a variance of **1.0**, but the variance of the new axis is **1.12**, which constitutes more than half of the total variance for the entire dataset (e.g., 1.12/2.00 or 56%).

Each value of theta will yield a different set of scores on Y1, and will also result in distinct values for the variance term. If we calculate transformed values and variances for different values of theta, we can compare the variance of the new axis to the total for our dataset.

Note that as we increase the

angle, the new variable accounts for an increasing fraction of total variance, until 45 degrees, and then declines; by the time theta is 90 degrees, the new axis is equivalent to X2, and, not surprisingly, its proportion of variance is back to 1.00 or 50.0%.

Theta	Var (Y1)	Proportion
5	1.121	56.0%
10	1.238	61.9%
15	1.348	67.4%
20	1.447	72.4%
25	1.533	76.7%
30	1.603	80.1%
35	1.654	82.7%
40	1.685	84.3%
45	1.696	84.8%
50	1.685	84.3%
55	1.654	82.7%
60	1.603	80.1%
65	1.533	76.7%
70	1.447	72.4%
75	1.348	67.4%
30	1.238	61.9%
35	1.121	56.0%
90	1.000	50.0%





Note that this time, the variance term is much smaller than that of the each of the original variables. But the variances of the two new axes sum to 2.0 -- the total variation in the original dataset. This is the basic approach of principal components analysis: obtaining linear combinations of variables into new axes, such that the first one accounts for the largest share of total variance, the second is orthogonal to the first and accounts for less variance, etc. Several properties hold for these components@JCurts/2009

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The transformation from (x1,x2) to (y1,y2) results several nice properties

- 1. The variability along y1 is largest.
- 2. Y1 and y2 are uncorrelated, that is, orthogonal.
- The confidence region based on (y1,y2) is easy to construct, and provide useful interpretations of the two sample plots.

Questions remain unanswered are

- 1. How to determine the angle θ so that the variability of observations along the y1 axis is maximized?
- 2. How to construct the ellipse for confidence region with different levels of confidences?
- 3. How to interpret the two-sample plots?

✓ How to determine the Y1 and Y2 axis so that the variability of observations along the Y1 axis is maximized and Y2 is orthogonal to Y1?

✓ Rewrite the linear relation between (Y1,Y2) and (x1,x2) in matrix notation:

 $Y1 = (\cos\theta) x1 + (\sin\theta) x2$ $Y2 = (\sin\theta) x1 + (\cos\theta) x2$

 $Y2 = (-\sin\theta) x1 + (\cos\theta) x2$

$$Y = \begin{bmatrix} y_1 \\ y_2 \end{bmatrix} = \begin{bmatrix} (\cos\theta)x_1 + (\sin\theta)x_2 \\ (-\sin\theta)x_1 + (\cos\theta)x_2 \end{bmatrix} = \begin{bmatrix} -\cos\theta & \sin\theta \\ \sin\theta & \cos\theta \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} = \begin{bmatrix} A_1'X \\ A_2'X \end{bmatrix} = AX$$

NOTE: X is bivariate , so is Y, and

V(X)

$$= \begin{bmatrix} V(x_1) & Cov(x_1, x_2) \\ Cov(x_1, x_2) & V(x_2) \end{bmatrix}, \quad V(Y) = A'V(X)A = \begin{bmatrix} V(x_1, x_2) \\ V(x_2) \end{bmatrix}$$

 λ_1 and λ_2 are called the eigen values. Which are the solutions of And, V(Y1) = λ_1 , V(Y2) = λ_2 , Correlation between Y1 and Y2 = 0.

$$\left|V(X) - \lambda I\right| = 0$$

 λ_1 and λ_2 are called the eigen values. Which are the solutions of And, V(Y1) = λ_1 , V(Y2) = λ_2 , Correlation between Y1 and Y2 = 0.



Note the angle depends on the correlation between X1 and X2, as well as, on the variances of X1 and X2, respectively.

• When ρ is close to zero, the angle is also close to zero. If V(X1) and V(X2) are close, then, the scatter plots are scattered like a circle. That is, there is no clear major principal component.

•When ρ is close to zero and V(X1) is much larger than V(X2), then, the angle will be close to zero, and the data points are likely to be parallel to the X-axis. On the other hand, if V(X1) is much smaller than V(X2), the angle will be close to 90^o, and the data points will be more likely parallel to the Y-axis. Consider, now, we actually observe the following two sample data:

 $\left\lceil \frac{1}{r} \right\rceil$

The sample means are given by
$$\begin{bmatrix} x_1 \\ \overline{x_2} \end{bmatrix}$$

The sample variance-covariance matrix is given
by
 $\hat{V}(X) = \begin{bmatrix} s_1^2 & rs_1s_2 \\ rs_1s_2 & s_2^2 \end{bmatrix}$
r is the Pearson's correlation coefficient, and S² is the
sample variance. S is the sample standard deviation.
 $\hat{V}(Y)$ is the solution of $\hat{V}(X) - \lambda I = 0$
The solutions for λ are given by $\underbrace{(s_1^2 + s_2^2) \pm \sqrt{(s_1^2 + s_2^2)^2 - 4 \cdot 1 - r^2)(s_1^2 s_2^2)^2}}_{(s_1^2 + s_2^2) \pm \sqrt{(s_1^2 + s_2^2)^2 - 4 \cdot 1 - r^2)(s_1^2 s_2^2)}}$

NOTE: V(Y1) + V(Y2) = $\lambda_1 + \lambda_2 = s_1^2 + s_2^2 = V(X1) + V(X2)$

Using the sample data, the angle is estimated by

$$\theta = (.5) \arctan\left(\frac{2rs_1s_2}{s_1^2 - s_2^2}\right)$$
$$= \arctan\left(\frac{\lambda_1 - s_1^2}{rs_1s_2}\right)$$

Correlation and Covariance Matrix



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Content Session - - X X **Eigen Analysis** Variable StDev Variance Ν N* Mean Χ1 25 10.880 4.503 20.277 0 COVA1 Analyze matrix: M1 CORR1 X2 10.680 4.905 24.060 25 M2 COVA1 Z1 25 0.000 1.000 1.000 0 Storage Z2 25 0.000 1.000 1.000 0 Column of eigenvalues: c11 Y1 with theta = 5 25 0 -7.155 3.909 15.277 -14.94 5.97 35.62 = 10 25 theta 0 Matrix of eigenvectors: m3 14.80 6.08 36.95 with theta = 4525 0 Y1 = theta = 9025 0 4.673 3.289 10.814 Boxplot of Stack data €. Help OK Cancel ACA data one.MTW *** 8 C5 C9 C11 C12 C 🔺 ŧ C1 C2 C3 C4 C10 C13 C14 X1 X2 Y1 with theta = 5 Y1 with theta = 10 Y1 with theta = 45 Y1 = theta = 90 Stack data Descriptor Z1 Z2 13 0.47080 -0.95411 -0.46098 6.0000 6 -2.0659-14,1721 11.9346 16 1 13 0.47080 0.67685 -9.7373 -18.5242 18,7418 6.69100 14.0000 1 17 14 13 0.47080 -10.6963 -19.0682 19.5927 15.0000 18 15 0.88072 7.58499 1 13 1.28846 -12.6141 -20.1563 21.2945 9.37299 17.0000 1 19 17 0.47080 -0.75024 7.0000 20 14 0.69288 -2.7412 -15.5551 13.3108 -0.01505 1 7 15 0.91495 0.47298 -8.2111 -19.6583 18.9416 4.90085 13.0000 21 13 1 22 17 13 1.35911 0.47298 -7.6438 -21.3365 19.9922 4.00471 13.0000 1 -11.4795 -23.5126 17.0000 23 17 17 1.35911 1.28846 23.3958 7.58069 1 1.69619 -13.1136 -25.4397 24 18 25.6230 8.92061 19.0000 1 19 1.58118 -13.5052 25 20 2.02533 1.90006 -27.6619 27.5245 8.91846 20.0000 1 20 4 7000 < ^^ 0 📳 Proje... 🗗 😐 🕺 @JCurt\$/2009 22 Calculate eigenvalues and eigenvectors for a symmetric matrix Editable

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Principal Component Analysis: X1, X2

Eigenanalysis of the Covariance Matrix

Eigenvalue 37.868 6.469 Proportion 0.854 0.146 Cumulative 0.854 1.000 Variable PC1 PC2 X1 0.663 0.748 X2 0.748 -0.663



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Principal Components Analysis - Storage																	
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pr _		Variables:		C4 Z2 C5 V1 with theta -	Scores:	c14c15											
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N		Number of components to	compute:		C10 Descriptor	Eigenvalues;	1010										
1.		Type of Matrix C Correlation			C12												
		Covariance			C13 C14												
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Ц-				a = 4	5 Y1 = theta = 90	Stack data	Descriptor										
1	-1.76959	-1.0669	-3.6053	3.2778	0.44377	2.0000	1	37.8677	0.663139	0.748496	3.4864	0.91921	37.8677				
2	-0.13863	-8.4546	-8.7965	10.6103	7.14767	10.0000	1	6.4690	0.748496	-0.663139	10.1375	-3.63741	6.4690				
3	-1.15798	-3.0926	-7.7545	7.4064	1.78154	5.0000	1				7.7213	1.17528					
4	-0.54637	-5.9694	-9.3866	9.9592	4.46353	8.0000	1				9.9668	-0.81414					
5	-0.13863	-7.8873	-10.4746	11.6610	6.25152	10.0000	1				11.4638	-2.14041					
6	-1.76959	0.0678	-6.9615	5.3791	-1.34852	2.0000	1				6.1390	3.91320					
7	0.47298	-10.4804	-12.9458	14.7390	8.48544	13.0000	1				14.3724	-3.38134					
8	-0.34250	-6.3610	-11.6088	11.8607	4.46138	9.0000	1				12.0416	0.01972					
9	-1.15798	-2.2417	-10.2717	8.9824	0.43732	5.0000	1				9.7107	3.42077					
10	-0.54637	-5.1184	-11.9038	11.5351	3.11931	8.0000	1				11.9562	1.43135					
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Pearson correlation of C14 and C15 = -0.000



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